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### Obtaining the Best Linear Unbiased Estimator of Noisy Signals by Non-Gaussian Component Analysis

Univers



M. Sugiyama<sup>1</sup>, M. Kawanabe<sup>2</sup>, G. Blanchard<sup>2</sup>, V. Spokoiny<sup>3</sup>, and K.-R. Müller<sup>2,4</sup>

Tokyo Institute of Technology, Japan
Fraunhofer FIRST.IDA, Germany
Weierstrass Institute & Humboldt University, Germany
University of Potsdam, Germany

### **Signal Denoising**

Signals we observe in practice are often noisy and redundant.



We want to remove noise  $\epsilon$  by cleverly making use of signal redundancy

## Setting

True signal is non-Gaussian:  $s \sim p(s)$ 

Noise is centered Gaussian:  $\epsilon \sim \phi(\epsilon)$ 

 $\blacksquare s$  and  $\epsilon$  are statistically independent .

We observe i.i.d. noisy samples:  $\{x_i\}_{i=1}^n$ 



Goal: obtain good estimates  $\{\widehat{s}_i\}_{i=1}^n$  of  $\{s_i\}_{i=1}^n$ 



- Projection does not have to be orthogonal.
- We want to choose "along"-subspace  $\mathcal{T}$  so that noise is maximally reduced.

# Best Linear Unbiased Estimator <sup>5</sup> (BLUE)

Linear estimator:  $\widehat{s}_i = Hx_i$ 

Unbiased estimator:  $\mathbb{E}_{\epsilon}[\widehat{s}_i] = s_i$ 

 $\mathbb{E}_{\epsilon}$ : Expectation over noise

BLUE: Minimum variance estimator among all linear unbiased estimators

 $\mathbb{E}_{\epsilon}(\widehat{s}_{i} - \mathbb{E}_{\epsilon}[\widehat{s}_{i}])^{2} \leq \mathbb{E}_{\epsilon}(\widetilde{s}_{i} - \mathbb{E}_{\epsilon}[\widetilde{s}_{i}])^{2}$ for any linear unbiased estimator  $\widetilde{s}_{i}$ 

### **Geometric View of BLUE**

Project 
$$\{x_i\}_{i=1}^n$$
 onto  $\mathcal{S}$  along  $\mathcal{T} = Q\mathcal{S}^{\perp}$ 



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### **Drawbacks of BLUE**

BLUE can be computed by

$$\widehat{\boldsymbol{s}}_i = \boldsymbol{H} \boldsymbol{x}_i$$
  $\boldsymbol{H} = (\boldsymbol{P} \boldsymbol{Q}^{-1} \boldsymbol{P})^{\dagger} \boldsymbol{Q}^{-1}$ 

†: Moore-Penrose

Thus we need generalized inverse (A) Noise covariance matrix Q (B) Projection matrix P (i.e., need to know S )

However, Q and P are often unknown.



### To Cope with (A)

Lemma: Q can be replaced with  $\Sigma = \mathbb{E}_{m{x}}[m{x}m{x}^{ op}]$  $H = (P\Sigma^{-1}P)^{\dagger}\Sigma^{-1}$ 

#### Intuition:

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- "*T*"-part of *Q* agrees with Σ since *s* ∈ *S* Utility: Estimating *Q* is not straightforward, but a consistent estimator of Σ can be constructed as

$$\widehat{\boldsymbol{\Sigma}} = rac{1}{n} \sum_{i=1}^{n} \boldsymbol{x}_i \boldsymbol{x}_i^{ op}$$

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### To Cope with (B)

Lemma: Non-Gaussian directions in signals is the signal subspace





# **Projection Pursuit**

(Friedman & Tukey, 1975)

 $\nu \sim \mathcal{N}(0,1)$ 

#### Iteratively finding non-Gaussian directions:

$$\widehat{\boldsymbol{\beta}} = \operatorname*{argmax}_{\|\boldsymbol{\beta}\|=1} \left| \mathbb{E}_{\boldsymbol{x}} [G(\boldsymbol{\beta}^{\top} \boldsymbol{x})] - \mathbb{E}_{\nu} [G(\nu)] \right|$$

G: projection index

• Kurtosis:  $G_1(\eta) = \eta^4$ (good for finding sub-Gaussians)

Gaussian mixture

Laplacian

• Robust index:  $G_2(\eta) = \frac{1}{b} \log \cosh(b\eta)$ (good for finding super-Gaussians) 10

# Multi-Index Projection Pursuit<sup>11</sup>

(Blanchard et al., 2006)

- Performance of PP depends on the choice of projection index.
- If samples contain both super- and sub-Gaussians, no single best index exists.
- MIPP combines results  $\{\widehat{\beta}_i\}$  of PP with many different indices by PCA.





### Normalization (cont.)

Such normalization is achieved by equalizing error orders of  $\{\widehat{\beta}_i\}$ :  $\varepsilon_i^2 = \mathbb{E}_x ||\widehat{\beta}_i - \mathbb{E}_x \widehat{\beta}_i||^2$ 



### Simulations



Super-Gaussian

0.16

0.14

0.12

0.1

0.08

0.06

0.04

0.02

Error

